Madras Agric, J. 71 (7) 431-438. July 1984.

CHARACTER ASSOCIATION AMONG COMPONENTS OF GENETIC VARIATION IN F1 GENERATION IN Arachis hypogea L.

V. ARUNACHALAM and A. BANDYOPADHYAYI

The nature and magnitude of association among 9 characters, 3 of which were measured in the early stage and the rest at harvest due to the components of variation—general combining ability, specific combining ability, reciprocal, maternal and non-maternal effects—was studied in two F1 diallel sets of crosses in groundnut. Pod yield, seed yield, 100 kernel weight and shelling per cent were positively and significantly correlated among themselves for all components of variation. Stepwise multiple regression analysis showed that pod yield, shelling per cent and 100-kernel weight were the major components to explain the variation in kernel yield; in addition, leaf area and specific leaf weight measured on 15-day old seedling were also found to be importment in one diallel. The favourable correlations among direct yield components have to be complemented by their desirable association with other characters spanning the growth phase of the plant, if direct selection for pod yield were to be successful. The contribution of reciprocal and maternal effects to the strength of character association was found to be significant.

Studies bearing on the association between quantitative characters including yield and its components are common in crop plants. A majority of investigations on groundnut reports positive and significant correlation of pod yield with number of pods, 100kernel weight and number of branches among others (Coffelt and Hammons, 1974: Dholaria et al., 1972: Hamid et al., 1981 Sadhu and Khera 1977. Sangha, 1973 and Singh et al. 1979). However, the material on which these results were obtained vary widely from bunch or spreading cultivars to Fr and F. of crosses between divergent parents,

One of the aims of a plant breeder is to understand the nature of association between important characters in a set of crosses. The information would be hiepful to identify a few 'key' char-

acters for which selection can fruitfully be made. If the crosses were made to a suitable design of mating like a full diallel, for instance, it would also be possible to partition the total covariation for any character into its various components - that due to general combining ability (gca), specific combining ability (sca) and reciprocal effects(rec) (Griffing, 1956). A study of the nature and magnitude of association among various characters for those components of variation would help to decide whether it would be possible to select parents or specific combinations for simultaneous improvement of important yield and other components of special interest to breeding for higher productivity. Results on such study are reported in this paper.

National Fellow, and Scientist-S 1, National Research Centre, IARI Regional Station, Rejendranager, Hyderabad 500 030, Andhra Predesh (INDIA) respectively.

MATERIAL AND METHODS

Two full diallel sets of crosses, one involving 15 parents (15-DL) and the other 10 parents (10-DL) were chosen for this study. The description of the parents, the characters studied on F₁ and details of field experimentation were described in an earlier paper (Arunachalam et al., 1982).

Analyses of variance and covariance were based on plot means and Griffings Model 1 Method 1 (Arunachalam et al., 1982; Griffing, 1956). The following 9 characters were considered:

FT-Days to first flowering;

LA-Leaf area of 15-day seedling;

SL-Specific leaf weight of 15-day seedling:

MP-Number of mature pods;

RP-Recovery percent [= 100 x no. of mature Pods/(no. of aerial pegs + no. of mature pods + no. of immature pods) taken at harvest]

SP-Shelling percent;

Tw-100-kernel weight;

PY-Pod yield; SY-Seed Yield.

The variation due to crosses was partitioned into various components-due to gca, sca and rec. The variation due to reciprocal effects was further partitioned into maternal (mat) and non-maternal (nmt) effects (Cockerham, 1963; Cockerham and Weir, 1977). Variance and covariance due to these sources were estimated and correlation coefficients between characters were computed for sources, gca, sca, rec, mat and nmt, in addition to the total genetic correlation coefficients [tgc).

The covariance matrices were further subjected to stepwise multiple, regression analysis (Draper and Smith 1966) keeping SY as dependent and others as independent variables. Variables were arranged in order of merit based on the proportion of variation in the kernel yield explained by them.

Correlation coefficients which were significant at 5 per cent in the desired direction were given a score +1; those which were significant in the undesirable direction were given a score -1; the non-significant ones received a score zero. For instance, early flowering and higher number of mature pods were desired. Hence a negative and significant correlation between these two characters would receive a score +1. High positive values were preferred for all characters other than FT. Thus negative and significant correlations with FT and in other cases, positive and significant correlations were given a score +1.

The possible correlation coefficients among 9 characters would be 36. The scores given to each of these correlation coefficients were added up and expressed as a proportion of the otal possible correlation coefficients—36in this case), to provide a correlation score. Thus the correlation score due o gca effects in 15-DL=3/36 (Table 2). In effect, the correlation score is ust the difference between the frequency of desirable and undesirable correlation coefficients, expressed as a proportion

It was possible to compute a correlation score for the possible correlation of any one subset of characters on any other. For instance, FT, LA and SL were characters measured in the early stage of crop growth. The rest of the characters were measured at harvest. The correlation between these two sets of characters can be judged by a correlation score computed over the 18 possible correlations (=1/18 for total genetic correlation in 15-DL. Table 2).

RESULTS AND DISCUSSION

Genetic correlation coefficients could be estimated between any two of the 9 characters in 15-DL. The estimate of genetic variance for flowering time was negative in 10-DL; hence only 28 correlation coefficients among 8 characters could be estimated (Table 1). Flowering time had significant and positive correlation with number of mature pods, recovery per cent and shelling per cent in 15-DL. Leaf area at seedling stage was however significantly and negatively correlated with number of mature pods and shelling per cent in both the diallels. Seedling vigour on the other hand had significant and positive correlation with recovery per cent in 15-DL and number of mature pods in 10-DL. The correlation score between characters measured in the early stage and at harvest was 1/18 in 15-DL and -2/12 in 10-DL both of which were low; in addition, the overall association between the two set of characters was nagative in 10-DL.

The correlation between any two of the major yield componentsood yield, seed yield. shelling per cent and 100-kernel weight-was significant and positive in both the diallels, for all components of variation, gca, sca, ec, mat, nmt and tgc.

A comparison of the nature of association can be made for various factors like gca, sca etc. in the two Fi diallels from Table 2. The correlation score for gca components was 5/15 (= 33 p. 100) in 10-DL as against 8/36 (= 88p. 100) in 15-DL. Similarly the correlation score for sca was also higher 10-DL than in 15-DL. Overall, the number of significant and positive correlations for reciprocal effects was 19 out of 36 in 15-DL and 13 out of 28 in 10-DL, both of the order of 50 p. 100 approximately. But the correlation score due to maternal effects was high in both the diallels, 30/36 (=83.p. 100) in 15-DL and 21/28 (=75p. 100) in 10-DL. The correlation score due to non-maternal effects was comparatively of a lower order (30 p. 100 approximately) (Table 2).

It may be noted that the correlation score computed for characters measured at harvest was positive for all components of variation in both the diallels. Further, every possible correlation coefficient among those characters for every component of variation was positive and significant in 10-DL while most of them was so in 15-DL, except for gca (Table 2).

The correlation score for characters measured at early stage and at harvest was negative only for gca and sca in 15-DL and for rec and nmt in 10-DL.

Pooling frequency of desirable and undesirable correlations over the two diallels, it was observed that the correlation score due to reciprocal effects was 32/64 (50 p. 100) and to maternal effects was 51/64 (80 p. 100).

The contribution due to non-maternal effects was 22/64 (= 34 p.:00) while that due to gca was 8/51 (16 p.:100) and sca 23/51 (45 p.:100).

The stepwise multiple regression analysis showed that pod yield was the major component to explain variation in kernel yield with respect to gca, sca, reciprocal, maternal or nonmaternal effects in both the diallels (Table 3). Next in importance to explain genetic variation in kernel vield were the characters, shelling percent, leaf area at seedling stage. 100-kernel weight and specific leaf area at seedling stage, (while flowering time was the least important) in 15-DL and shelling per cent and 100-kernel weight in 10-DL.

Leaf area was also found to be an important component next to pod yield in explaining the variation in kernel yield due to non-maternal effects in 15-DL, and due to sca in 10-DL.

Inferences drawn from correlation studies are influenced to a large extent by the number of genotypes on which they are based. Segregating genotypes, fewer number of characters and their nature (whether continuously varying or not) will also impose restraint on the validity of inferences. Further it is not uncommon to come across results on correlation based on plot or plant means when the material was raised in field designs like randomised blocks or split-plot. In such cases, the environmental variation estimated by the field design does not get accounted for in the estimates of correlation coefficients. In this study, therefore, the various components of variation like gce, sca,

rec, mat, nmt and tgc have been estimated using the relevant genetic model
and the randomised blocks design on
which the F₁ diallels were evaluated.
The parents of the two diallels were also
represented by cultivars from all the
botanical sub-groups, Spanish Valencia
and Virginia (Arunachalam, 1976). The
results on character association based
on the diallels would thus be broadbased.

Positive and significant correlation were observed among the four yield components, pod and seed yield, 100kernal weight and shelling percent for all the components of genetic variation. The results concur with the results reported by Hamid et al., (1981) and Coffelt and Hammons (1974) where the correlation among those characters was studied on six crosses in F, generation in the former and on six population including parents and Fr of direct and reciprocal crosses in the latter. It should be noted that the results of Hamid et al., (1981) were based on segregating genotypes in F. from different crosses. Further, gca effects were also estimated from Fa data. Since different samples of F. could contain different genotypes also varying in number, inferences drawn on correlations based on gca effects estimated from F, were of limited value. However, in their study the variation due to reciprocal differences was not significant. But, in this study, significant correlation among characters due to reciprocal and maternal effects were observed in the F1 of both diallels. The correlation score due to gca was much less 'than that due to sca, pointing to the importance of non-additive gene action in bringing

Table 1: Correlation among 9 characters due to various genetic components of variation

		gca -	s	ca	t	90	π	at	n;	mt	.1	gc
	а	b	a	b	8	ь	8	ь	а	b	a	b
FT-LA	-1	0	-1	0	-1		i		~f	====	-1	
FT-SL	1	0	3		4		4		1		1	_
FT-MP	1		0	-1	1		1		0		1	_
FT-RP	- 1		1	13	1		1	-	1	_	1	- 2
FT-SP	1	1	0	-	1		1		0	2	1	_
FT-TW	0	_	0	1	0	-	1		0	-	0	- 2
FT-PY	-1	1	0		1		3	42.7	1		0	_
FT—SY	-1	10	0	1	1		4		0		0	- 1
LF-31.	1	0	1	-	ō	1	1	0	-1	0	0	0
LAIAF	1-1	· <u>-</u> -	0	0	-1	0	1	1	1	0	1	-1
LA-RP	-1	-	-1	0	-1	1	1	1	-1	1	-1	0
LA-SP	-1	-1	-1		0	1	1	1	-1	-1	1	-1
LA-TW	0		1	1	1	-1	1	-1	0	0	0	-1
LA-PY	0	0	0	ÿ. ;	-1	-1	-1	-1	-1	-1	0	0
LA -SY	-1	0	0	4	1	-4	-1	-1	-1	-1	0	0
SL-MF	0	1,000	1	· -	1	1	0	-1	1	0	0	Ĩ
SL -RP	1	-	1	4	1	0	1	1	1	0	1	0
SL -SP	. 1	0	-1		1,	1	t:	:1	1	-1	0	0
SL -TW	0	: 22	-1	1-1	0	0	1	1	0	-1	0	ō
SL -PY	-1	Ö	-1	1 2	1.	-1	1	1	1	1	0	0
SL -SY	-1	0	-1		1	-1	1	1	1	-1	0	0
MP -RF	1	1-0	1	1	1	1	1	1	1	1	1	1
MP -SF	3		1	-	0	1	0	1	0	î	- 1	1
MP-TV	-1		0	1	-1	1	-1	1	0	1	-1	0
MP -PY	1	1	1	*,	- 1	1	1	1	0	1	-1	0
MP -SY	-1	1	1	1	1	1	1	1	- 1	1	-1	1
RP -SP	11:	. <u>***</u> *	1	-	1.	1	1	1	1	1	0	1
RP -TW	-1	77	1	1	1	1	1	1	-1	1	0	0
RP -PY	1	1 22	1	-	-1	1	1	1	1	1	0	1
RP -SY	-1		1	1	1	1	1	1	.1	1	1	1
SP _TW	1		1	+	4	1	1	7	1	1	1	1
SP -PY	1	1	1	===	77	1	1	1	1	1	1	3
SP-SY	1	1	1	, n 40 1	1	1	1	1	1	1.	1	1
TW -PY	1	-	1	77	1	1	1	1	1	1	1	1
TW-SY	1	-	1	1	-1	1. 1	:1	1:	1	*	1	1
PY-SY	1	1	1	_	1	1	1	1	1	1	1	1

a=15-DL; b=10-DL I +1=Significant correlation in the desirable direction;

⁻¹⁻Significant correlation in the undesirable : 0-Non-significant - = Nonestimable : For other symbols, see text.

Table 2. Correlation scores among characters measured at early stage (E), harvest (Y)

E vs Y and overall characters

		E	Y	E vs Y	Overall
15-DL	g	1/3	3/15	-1/18	3/36
-3,-	s	1/3	14/15	-2/18	13/36
	r	0]3	12/15	7/18 -	19/36
	m	3/3	12/15	15/18	30/36
	nm	-1/3	12/15	2/18	13/36
	t.	0/3	6/15	1/18	7/36
10-DL	g	0/3	3/3	2/9	5715
	s	0/1	6/6	4/8	10/15
	1	1/1	15/15	-3/12	13/28
	m	0/1	15/15	6/12	21/28
	nm	0/1	15/15	-6/12	9/28
	t	0/1	12/15	-2/12	10/28
Overall	· 9	1/6	6/18	1/27	8/51
	5	1/4	20/12	2/26	23/51
	i i	1/4	27/30	4/30	32/64
	m	3/4	27/30	21/30	51/64
	nm	-1/4	27/30	-4/30	22/64
	t	0/4	18/30	-1/30	17/64

about positive and significant character association.

If the possible character correlations were classified as those among
early stage characters, among yield
components and one vs. the other
(Table 2), it was seen that the high
frequency of positive and significant
associations among yield components
for gca and sca was not directly reflected in the frequency of correlations
among all the characters (given under
the column 'overall' in Table 2) mainly
due to the low frequency of desirable
(or occurrence of undesirable) correlations between early stage characters
and yield components. On the contrary,

the frequency of those correlations for reciprocal and maternal effects were high, positive and significant resulting in high 'overall' correlation scores (Table 2).

The results suggested that selection for pod yield alone can be as effective as simultaneous selection for other yield components like seed yield. 100-kernel weight and shelling percent.

Other than pod yield, the two early stage characters, leaf area and specific leaf weight at seedling stage were found to be important in explaining

Table 3 Order of importance of various characters in explaining the variation in kernel yield in two F, diallels

	Genetic			gca		ě	sca			rec		-	mat		1	umt	
8	8	M	0	В	MC	8	æ	MC OV	8	æ	MC	8	В	MC	0	00 В	MC
16.01	71												÷				
β	0,5490		Ā	0.9538 PY 0.591m 0.8708	0.8708	7	PY 0,440 _m		γ	0,9311 PY 0,598 m	0.9338	à.	0.9338 PY 0.725∞		PY	0.9761 PY 0.581a	0.961
SP	0,6470	0.9879				SP	0.454°	0 9879	Ā	PY 0.55300	0.9815				Z	LA -0.407m	0.9853
LA.	LA- 0.22100	0.9896				≥	TW 0,399®	0.9975	RP.	RP -0.2760	0.9846						
≥	TW 0 095m	0,9902															
SL	SL 7.418	0.9904															
M	MP 0,22800	0 9905															
R.P	RP -0,34100	0,9914															
t	0,001	0.9920													ï		
10.01	70		,														
ΡY	0.633∞	0,9766	ď	PY 0,493 0,9830	0.9830	숣	PY 2,447∞	0.9215	ě	1,768∞	0.9215 PY 1,768 0,9786	ď	PV 0,5490 0,3942 TW 1 8520	0,394	ř.	1 8528	0 999
SP	0.4310	0.9910		TW 0,66800	0.9998	ΓĄ	LA -0,79300	0.9888									
3	TW -0.514 to 0.9972	0,9972	, in the			1					1						

OV - Order of variable; B - fegfession coefficient; MC - multiple correlation coefficients R4; P - Significant at 5 p. 100; for other symbols, see text. for other symbols, see text, Significant at 5 p. 100;

the variation for kernel weight (Table3). This observation confirmed the need for yield components to have desirable association with early stage and on the same argument, other important characters in the entire growth phase of the plant, if selection for economic characters were to be successfull.

The results on character association (Table 2) have substantiated the need to select not only parents but also to decide whether they should be used as a male or female in specific crosses. Further, the difference in the overall correlation score in the two diallels (7/36 in 15-DL vs. 10/28 in 10-DL) brought to focus the fact that character association was highly sensitive to the nature and frequency of entries on which its was based and has limited scope for generalisation.

REFERENCES

- ARUNACHALAM, V. 1976. The utility of covariance of combining ability in plant breeding. Theor. appl. Genet. 47: 303-06.
- ARUNACHALAM, V., A. BANDYOBADHYAY'
 S. N. NIGAM and R. W. GIBBONS. 1982.
 Heterotic potential of single crosses in
 grounJnut (Arachis hypogaea L.). Oleagineux, 37: 416-420.
- COCKERHAM, C. C. 1963. Estimation of genetic variances. 53-93. In: Statistical Genetics and Plant Breeding. NAS-NRC Pub. 982, Washington, U. S. A.
- COCKERHAM, C. C. and B. S. WEIR, 1977.

 Quadratic analyses of reciprocal crosses
 Biometrics, 33: 187-203.
- COFFELT, T. A. and R. O. HOMMONS. 1974.

 Correlation and heritability studies of nine characters in parential and intra-specific populations of Arachis hypogaea. Oleagineux, 29: 23-7.
- DHOLARIA, S. J., S. N. JOSHI and M. M. KABARIA. 1972. Correlation of yield and contributory characters in groundnut grown under high and low fertility levels. *Indian J. Agric. Sci.*, 42 1084-86.

- DRAPER, N. R. and H. SMITH. 1966. Selection of the best regression equation. 163-195. In: Applied Regression Analysis. John Wiley & Sons Inc., New York, U. S. A.
- GRIFFING, B. 1956. Concept of general and specific combining ability in relation to diallel crossing systems. Aust. Jour. Biol. Sci., 9: 463-93.
- HAMID, M. A., T. G. ISLEIB, J. C. WYNN and C. C. GREEN. 1981. Combining ability analysis of cercospora leaf spot resistance and agronomic traits in *Arachis hypogaea* L. *Oleagineux*, 36: 605-12.
- SANDHU, B. S. and A. S. KHERA, 1977. Interrelationship in semi-spreadin bunch and semi-spreading crosses of groundnut. Indian J. Genet., 37: 22-6.
- SANGHA, A. S. 1973. Genetic diversity in spreading groundnut. *Madras Agric. J*, 60: 1380-87.
- SINGH, A. S., M. SINGH and K. S. LABANA, 1979, Variability and correlation studies in groundnut after hybridisatiod. *Madras Agric. J.*, 66: 585-70.